

**APPENDIX A**  
**EUC PROVISIONS, BY PHASE**



TABLE A.1

## SUMMARY OF THE FIVE PHASES OF EUC

EUC Phase	Public Law	Date Enacted	Dates of Eligibility		Duration and Trigger Levels	EB	Funding	Options to Claim	Other
			Effective	Termination					
1	102-164 (the Emergency Unemployment Compensation Act of 1991)	November 15, 1991	November 17, 1991  Reachback provisions for those whose benefit year ended after February 28, 1991	July 4, 1992	Created three tiers of EUC benefit durations, at 6, 13, or 20 weeks. States had 13 weeks if the AIUR was at least 4 percent in the preceding 12 weeks and the current week or if the AIUR was 2.5 percent and the exhaustion rate 29 percent. States had 20 weeks if the AIUR was at least 5 percent in the preceding 12 weeks and the current week or if the average TUR was at least 9 percent in the previous 6 months.	Allowed governors to deactivate EB to pay EUC. In states in which EB was in use, claimants received EUC benefits only after EB benefits were exhausted and only the amount in excess of the amount paid through EB.	Funds in the Extended Unemployment Compensation Account (EUCA) were used. A one-year extension of the 0.2 percent Federal Unemployment Tax Act surtax and a variety of offsets and tax extensions were used to meet Budget Enforcement Act (BEA) requirements.		
	102-182	December 4, 1991	Retroactive to November 17, 1991	June 13, 1992	Eliminated the 6-week tier; for those states, individuals were eligible for 13 weeks.				
2	102-244	February 7, 1992	February 8, 1992	July 4, 1992	Weeks of maximum benefits were increased to either 26 or 33 weeks, subject to a maximum of 130 percent of a claimant's regular benefit payments, for claimants filing initial claims from February 9, 1992, to June 13, 1992. For claimants filing after June 13, 1992, or for claimants with nonconsecutive weeks claimed, the maximum benefits were decreased to either 13 or 20 weeks.		BEA funding requirements were met via a temporary acceleration in payments of corporate income taxes and carryover pay-as-you-go financing from earlier entitlement legislation.		

TABLE A.1 (continued)

EUC Phase	Public Law	Date Enacted	Dates of Eligibility		Duration and Trigger Levels	EB	Funding	Options to Claim	Other
			Effective	Termination					
3	102-318	July 3, 1992	Retroactive to June 13, 1992	No new claims after March 6, 1993. No payments after June 19, 1993.	<p>Set EUC durations to either 20 or 26 weeks, subject to a maximum of 100 percent of a claimant's regular benefit payments for the most recent benefit year.</p> <p>Benefit durations dropped to either 10 or 15 weeks, subject to a maximum of 60 percent of a claimant's regular benefits, depending on the state unemployment rate when the seasonally adjusted national TUR was 6.8-7 percent for 2 consecutive months.</p> <p>Benefit durations dropped to either 7 or 13 weeks, subject to a maximum of 60 percent of a claimant's regular benefits, depending on the state unemployment rate when the seasonally adjusted national TUR was less than 6.8 percent for 2 consecutive months.</p>	<p>Governors' option to deactivate EB to pay EUC was not applicable for any EB period beginning after March 6, 1993. New claimants in states that used EB and had an EUC balance after March 6, 1993, could receive payments from the program (EB or EUC) with the greater balance.</p> <p>Allowance for usage of 3-month average TUR as trigger for EB. Changed EB durations from exactly 13 to either 13 or 20 weeks.</p>	All EUC benefits funded by general revenue funds deposited in the EUCA, satisfying BEA requirements through several non-UI income and tax adjustments and carryover pay-as-you-go financing from earlier legislation.	An individual could defer rights to regular UI benefits for weeks of unemployment beginning on or after July 3, 1992, to collect EUC associated with the most recent prior benefit year.	States could use more than one method to measure employment and earnings (had at least 20 weeks of employment in the base period, earned 150 percent of the base period high quarter wages during the base period, or earned wages during the base period of at least 40 times the claimant's weekly benefit amount for regular benefits) for qualifying purposes, rather than using one method exclusively, as was required prior to July 3, 1992.

TABLE A.1 (continued)

EUC Phase	Public Law	Date Enacted	Dates of Eligibility		Duration and Trigger Levels	EB	Funding	Options to Claim	Other
			Effective	Termination					
4	103-6	March 4, 1993	March 6, 1993	No new claims after October 2, 1993. No payments after January 15, 1994.		Governors' option to deactivate EB to pay EUC was not applicable for any EB period beginning after October 2, 1993. New claimants in states that used EB and had an EUC balance after October 2, 1993, received payments in the program (EB or EUC) with the greater balance.	All EUC benefits for initial claims attributable to weeks of unemployment beginning after October 2, 1992, funded by general revenue funds included in the DOL Appropriations Acts and then transferred to the EUCA.		Changed the work search requirements from those in the EB provisions to those in state law provisions for regular UI.
	103-6	July 26, 1993			Revised interpretation of the 7 percent and 6.8 percent thresholds of the national TUR before EUC durations changed (per P.L. 102-318). The 7 percent period would be in effect when the national TUR for <i>each of the 2 most recent months</i> was less than 7 percent, rather than when the average of the 2 months was less than 7 percent. Similar interpretation for the 6.8 percent threshold.				

TABLE A.1 (continued)

EUC Phase	Public Law	Date Enacted	Dates of Eligibility		Duration and Trigger Levels	EB	Funding	Options to Claim	Other
			Effective	Termination					
5	103-152	November 24, 1993	Retroactive to October 2, 1993	No new claims after February 5, 1994. No payments after April 30, 1994.	Amended maximum number of benefit weeks to either 7 or 13, subject to a maximum of 50 percent of the claimant's regular benefits.	Governors' option to deactivate EB to pay EUC was not applicable for any EB period beginning after February 5, 1994. New claimants in states that used EB and had an EUC balance after February 5, 1994, received payments in the program (EB or EUC) with the greater balance.	Benefits for initial claims-attributable to unemployment beginning after October 2, 1993, were paid from the EUCA, financed through savings from profiling requirements, elimination of choice in filing, and increases in the sponsor-to-alien deeming period under Supplemental Security Income.	Repealed the option established in P.L. 102-318, whereby a claimant could choose either to file a new claim or receive EUC on the basis of a prior benefit year.	

TABLE A.2

## DURATIONS OF EUC BENEFITS OVER TIME, BY STATE, IN WEEKS

	EUC-1 P.L. 102- 162 and 102-182 11/17/91	State- Specific Duration Changes While EUC- 1 in Effect	EUC-2 P.L. 102-244 2/8/92	State-Specific Duration Changes While EUC-2 in Effect	EUC-3 P.L. 102-318 6/14/92	State-Specific Duration Changes While EUC-3 in Effect	Trigger Changes from EUC-3 and EUC-4 P.L. 103-6 3/6/93	State-Specific Duration Changes While EUC-4 in Effect	EUC-5 P.L. 103- 152 10/2/93	State-Specific Duration Changes While EUC-5 in Effect
AL	13		26		20		10		7	
AK	20		33		26		15		13	1/23/94--onto EB
AZ	13		26		20		10		7	
AR	13	2/2/92--20	33		20		10		7	
CA	13	1/5/92--20	33		26		15		13	
CO	13		26		20		10		7	
CT	20		33		26	11/1/92--20	10		7	
DE	13		26		20		10		7	
DC	13		26		20		10		7	
FL	13		26		20		10		7	
GA	13		26		20		10		7	
HI	13		26		20		10		7	
ID	13	2/9/92--20	33		26	7/19/92--20 2/21/93--26	15	7/4/93--10	7	
IL	13		26		20		10		7	
IN	13		26		20		10		7	
IA	13		26		20		10		7	

TABLE A.2 (continued)

	EUC-1 P.L. 102- 162 and 102-182 11/17/91	State- Specific Duration Changes While EUC- 1 in Effect	EUC-2 P.L. 102-244 2/8/92	State-Specific Duration Changes While EUC-2 in Effect	EUC-3 P.L. 102-318 6/14/92	State-Specific Duration Changes While EUC-3 in Effect	Trigger Changes from EUC-3 and EUC-4 P.L. 103-6 3/6/93	State-Specific Duration Changes While EUC-4 in Effect	EUC-5 P.L. 103- 152 10/2/93	State-Specific Duration Changes While EUC-5 in Effect
KS	13		26		20		10		7	
KY	13		26		20		10		7	
LA	13		26		20		10		7	
ME	20		33		26	8/30/92--20	10	3/28/93--15 6/27/93--10	7	3/27/94--onto 20 weeks EB
MD	13		26		20		10		7	
MA	20		33		26	8/2/92--20	10		7	
MI	20		33		26	10/25/92--20	10		7	
MN	13		26		20		10		7	
MS	20		33	2/16/92--26	20		10		7	
MO	13		26		20		10		7	
MT	13		26	3/8/92--33	20		10	3/7/93--15 6/12/93--10	7	
NE	13		26		20		10		7	
NV	13		26	3/8/92--33 6/6/92--26	20		10		7	
NH	13		26		20		10		7	
NJ	20		33		26	11/22/92--20	10	3/7/9--15 6/13/93--10	7	
NM	13		26		20		10		7	
NY	13		26	2/16/92--33	26	7/12/92--20	10		7	



TABLE A.2 (continued)

	EUC-1 P.L. 102- 162 and 102-182 11/17/91	State- Specific Duration Changes While EUC- 1 in Effect	EUC-2 P.L. 102-244 2/8/92	State-Specific Duration Changes While EUC-2 in Effect	EUC-3 P.L. 102-318 6/14/92	State-Specific Duration Changes While EUC-3 in Effect	Trigger Changes from EUC-3 and EUC-4 P.L. 103-6 3/6/93	State-Specific Duration Changes While EUC-4 in Effect	EUC-5 P.L. 103- 152 10/2/93	State-Specific Duration Changes While EUC-5 in Effect
NC	13		26		20		10		7	
ND	13		26		20		10		7	
OH	13		26		20		10		7	
OK	13		26		20		10		7	
OR	13	1/12/92--20	33		26	9/27/92--20 1/31/93--26	15	7/11/93--10	7	10/3/93--onto EB 2/26/94--off EB
PA	13	1/26/92--20	33		26	8/16/92--20	10	3/21/93--15 6/20/93--10	7	
RI	20		33		26		15		7	1/16/94--13
SC	13		26		20		10		7	
SD	13		26		20		10		7	
TN	13		26		20		10		7	
TX	13		26		20		10		7	
UT	13		26		20		10		7	
VT	13	1/19/92--20	33		26	8/16/92--20	10	5/09/93--15 8/8/93--10	7	
VA	13		26		20		10		7	
WA	13	2/2/92--20	33		26	7/4/92--20 1/31/93--26	15	6/27/93--10	7	10/3/93--onto EB 2/26/94--off EB

TABLE A.2 (continued)

	EUC-1 P.L. 102- 162 and 102-182 11/17/91	State- Specific Duration Changes While EUC- 1 in Effect	EUC-2 P.L. 102-244 2/8/92	State-Specific Duration Changes While EUC-2 in Effect	EUC-3 P.L. 102-318 6/14/92	State-Specific Duration Changes While EUC-3 in Effect	Trigger Changes from EUC-3 and EUC-4 P.L. 103-6 3/6/93	State-Specific Duration Changes While EUC-4 in Effect	EUC-5 P.L. 103- 152 10/2/93	State-Specific Duration Changes While EUC-5 in Effect
WV	20		33		26		15		13	
WI	13		26		20		10		7	
WY	13		26		20		10		7	

SOURCE: Unpublished table "Emergency Unemployment Compensation Periods," by U.S. Department of Labor and Federal Register, Washington, DC, U.S. Government Printing Office, various days.

**APPENDIX B**

**SAMPLE DESIGN AND SAMPLE WEIGHTS**

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The sample for the Emergency Unemployment Compensation (EUC) evaluation was designed to represent the national population of EUC recipients and to provide sufficient statistical precision to meet the descriptive and analytic objectives of the study. It was also designed to provide a comparison group of Unemployment Insurance (UI) recipients who did not receive EUC; this group was representative of the national population of UI-only recipients when EUC was available. More specifically, the sample design called for a two-stage sampling process: initially, 23 states were selected; then, recipients in those states were selected. Administrative records were to be collected and analyzed for the recipient samples and survey data were to be collected for subsamples.

In practice, a number of states selected for the sample were unable to participate. Additional states were selected and asked to participate, but, in the end, only 18 of the 35 states that were asked provided samples of recipients. In addition, response rates for the survey were low (just under 50 percent), primarily due to difficulty in locating respondents (see Appendix C).

Both state and respondent nonresponse raise the possibility that estimates from the samples may be biased. However, our analysis of this issue suggests that the administrative records samples from the 18 states can be weighted to represent the national population on key dimensions of UI receipt. Therefore, we believe that the results we obtain with these samples can be characterized as representing the nation. We use an analogous procedure to weight the 16 state survey samples to be nationally representative.<sup>1</sup> Furthermore, our analysis of survey nonresponse (Appendix C) suggests that the respondents are similar to nonrespondents on key demographic and UI receipt characteristics.

We now turn to a discussion of the sample design and our procedure for computing weights.

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<sup>1</sup>Administrative samples from two states were received too late to be included in the survey.

## **A. INITIAL SAMPLE DESIGN**

The sample design for the EUC evaluation was intended to fulfill three main objectives. First, it was designed to produce a sample that was representative of the national population of EUC recipients. Second, it was intended to provide a comparison group of UI recipients who did not receive EUC that was representative of the national population of UI-only recipients when EUC was available. Third, it was meant to provide sufficient statistical precision for the descriptive and analytic objectives of the study.

To address the first objective, we defined the EUC sample frame as all individuals in the 51 states who received an EUC payment.<sup>2</sup> We planned to select a sample from this sample frame and to collect administrative records data for this sample. We also planned to collect survey data for a subsample but to limit the survey subsample to individuals who began collecting EUC in July 1992 or later. We restricted the survey subsample because we wanted to limit the period for which recipients were asked to recall labor market events. We chose July 1992, which was the start date of EUC Phase 3, so that the survey sample would be representative of EUC recipients in Phases 3 through 5.

To address the second objective, we defined the UI-only sample frame as all individuals in the 51 states who began collecting UI between January 1991 and September 1993 and who did not collect EUC. We chose these start and end dates for this sample to capture the majority of UI recipients who could have collected EUC. Although some individuals who began collecting UI as early as March 1990 collected EUC through its reachback provisions, the number of such individuals was small relative to the entire UI population. For this reason, we restricted the comparison group

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<sup>2</sup>We included in our universe the 50 states plus the District of Columbia. For convenience, we refer to this group as the "51 states."

to individuals who were more likely to transition to EUC if they exhausted UI. Individuals who began collecting UI in early 1991 would have exhausted UI in the second half of 1991 and could have collected EUC beginning in November 1991. We chose September 1993 as the end date for this comparison sample for similar reasons. Some individuals who started collecting UI after September could have exhausted UI and begun collecting EUC prior to February 5, when the last EUC initial claims were taken, but most individuals who ended up on EUC would have begun collecting UI earlier. Finally, we decided that the UI-only interview subsample would include UI-only recipients who began collecting UI between January 1992 and September 1993. We chose January 1992 as the start date to include individuals who would have been likely to collect EUC beginning in July 1992 or later if they had collected EUC.

To address the third objective, we decided that a reasonable precision standard for the survey subsamples would involve describing attributes of the EUC population with a  $\pm 2.5$  percent, 95 percent confidence interval and differences between the EUC and UI-only samples of  $\pm 6.0$  percent at 95 percent confidence, for attributes with an incidence of 50 percent in the population. We calculated that these objectives could be achieved with roughly 1,500 EUC and 900 UI-only sample members, if the samples were simple random samples of the national population.<sup>3</sup>

Because the UC program operates separately in each state, however, it was, not feasible to select simple random samples from the national population of EUC and UI-only recipients. Instead, we chose a two-stage sampling procedure that involved the random selection of states in the first stage and recipients in the second stage. Specifically, we decided to choose states in the first stage with probability proportional to the size of their EUC population and then to choose equal-sized samples

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<sup>3</sup>We used a two-tail test at the 80 percent power level for this computation.

of EUC recipients in the second stage.<sup>4</sup> This procedure maintained equal probabilities of selection for all EUC recipients and was intended to yield a self-weighting sample of EUC recipients. A comparable UI-only sample was allocated to each state in a way that was designed to provide a self-weighting sample of such individuals.<sup>5</sup>

Because of the two-stage sample design, we also had to increase the EUC and UI-only sample sizes to take into account the loss of statistical precision (termed the "design effect") resulting from clustering the sample in a limited number of states. To account for the importance of design effects, we considered the degree to which average UI benefit duration varies across states.<sup>6</sup> In 1991,

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<sup>4</sup>Since the EUC caseload was heavily concentrated in a few states, this procedure was modified slightly to allow for the fact that the sample would definitely contain the largest states. Once these states were identified, sample sizes were allocated to them in proportion to their representation in the national caseload. The remaining states were then selected with probabilities proportional to size, with equal size samples being allocated to each state.

<sup>5</sup>To draw a nationally representative sample of regular UI-only recipients, we needed to account for the fact that the selection probabilities of states were relative to the EUC population, rather than to regular UI-only recipients or to recipients in general. Following the approach used in an earlier study, regular UI-only recipients were sampled with equal probabilities of selection by allocating larger numbers of regular UI-only recipients to states with smaller numbers of EUC recipients, according to the following formula (Corson and Dynarski 1990):

$$(1) \quad Q_j = X_j [ (1 - E_j) / E_j ] R.$$

where, for state  $j$ ,  $Q_j$  is the regular UI-only sample,  $X_j$  is the expected size of the subsample of our sample of EUC claimants who collected regular UI earlier in their unemployment spells,  $E_j$  is the ratio of the total number of EUC recipients who previously collected UI in the state to the total number of UI recipients in the state, and  $R$  is the uniform sampling rate required to adjust the size of the UI-only sample to the desired total number.

<sup>6</sup>We used average benefit duration for regular UI to assess the importance of design effects. Although other variables would yield different results, we expected that the variation among states on this variable would indicate variation in important outcome variables, such as duration of EUC receipt.



average UI duration nationwide was 15.8 weeks; however, an examination of average duration by state revealed important systematic variation. Average duration was more than 17 weeks in five states and less than 12 weeks in nine states. Because earlier studies of the UI population (see, for example, Corson and Dynarski 1990) suggested that the total variance in average UI duration is about 144 weeks, we used the variation in state-level averages to estimate the state component of variance and allocated total variance between individuals and states as follows:<sup>7</sup>

Variance Component	Variance	Percentage
Individual Recipient	137.6	95.6
State	6.4	4.4
<b>Total</b>	<b>144.0</b>	<b>100.0</b>

These data suggested that 4.4 percent of the variability in average benefit duration is attributable to state-specific factors and the remaining 95.6 percent to recipient-specific factors. Although 4.4 percent at first seems like a small amount, it is a major component of variability for a sample of EUC recipients drawn from a small subset of states.

We explored the implications of this situation for various recipient and state sample sizes. We found, for example, that the standard deviation of the estimate of average benefit duration made from a simple random sample of 2,500 recipients drawn from all 51 states would be .24 (the "one-stage" simple random sample estimate). If the sample was restricted to 15 states, the standard deviation would be .42, a difference of 75 percent. In this example, the sample of 2,500 recipients drawn from

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<sup>7</sup>The weighted state-level variance in average duration is equal to  $\sum w_s(d_s - d_n)^2$ , where  $w_s$  is the state share of the population, and  $d_s$  and  $d_n$  are state average duration and national average duration, respectively.

15 states would provide the same statistical precision as a one-stage simple random sample of only 821 recipients (the "effective" sample size) drawn from all 51 states. Increasing the recipient sample size would do little to improve precision, because the source of the high variance is state specific, not recipient-specific. For example, doubling the sample to 5,000 recipients drawn from the same 15 states would only increase the effective sample size from 821 to 974 (an increase of 19 percent). Instead, substantially greater gains in precision could be achieved by increasing the number of states. For example, with 23 states instead of 15, a sample of 2,500 recipients has an effective size of 1,547, compared with 821 for 15 states.

On the basis of this analysis, we decided to draw our sample from 23 states and to interview 2,500 EUC recipients and 1,500 UI-only recipients. Eleven states (New York, California, Pennsylvania, Texas, New Jersey, Illinois, Florida, Michigan, North Carolina, Ohio, and Massachusetts), representing 64 percent of the EUC population, were selected with certainty and allocated 64 percent of our sample (see Table B.1).

The remaining noncertainty states could have been selected by a simple random drawing from the remaining states with probabilities of selection proportional to size; however, we believed additional stratification was warranted. Specifically, we chose the 12 noncertainty states on the basis of a stratified sample according to average UI benefit duration.<sup>8</sup> This stratification was intended to ensure adequate variability in the sample along dimensions, such as labor market strength and generosity of state UI, programs that are approximated by the average duration figures. To accomplish the stratification, the 42 noncertainty states were grouped into three equal-sized strata--high, medium and low duration--with four states being selected from each stratum as shown in Table B.1.

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<sup>8</sup>To ensure regional representativeness, we ordered states within stratum by region.

TABLE B.1  
STATE SELECTION PROBABILITIES

Selection Criteria	State	DOL Region	Number of EUC First Claims	Average Benefit Duration	Selection Probability	Supplementary Sample	
					23 States	States in Initial Sample	States in Supplementary Sample
<b>Certainty States</b>	NY	2	1,099,894	20.03	1	X	
	CA	9	1,030,755	16.54	1	X	
	PA	3	594,664	16.70	1	X	
	TX	6	528,744	15.17	1	X	
	NJ	2	479,865	18.05	1	X	
	IL	5	466,784	17.23	1	X	
	FL	4	464,163	15.02	1	X	
	MI	5	422,678	14.51	1	X	
	NC	4	322,288	10.50	1	X	
<b>High-Duration States</b>	OH	5	272,271	14.71	1	X	
	MA	1	252,241	18.95	1	X	
	ME	1	81,584	15.84	0.30	X	
	VT	1	20,676	16.14	0.08		
	CT	1	198,648	16.19	0.73		X
	RI	1	83,076	16.73	0.30	X	
	WV	3	55,519	15.12	0.20		X
	MD	3	140,084	16.69	0.51		
	DC	3	44,254	20.51	0.16	X	
<b>Medium-Duration States</b>	MN	5	110,940	15.65	0.41		X
	NM	6	14,854	15.89	0.05		
	OR	10	129,269	14.93	0.47	X	
	AK	10	43,790	15.33	0.16		
	WA	10	177,344	16.10	0.65		X
	DE	3	15,694	14.22	0.06		
	MS	4	85,884	13.23	0.31	X	
	KY	4	90,465	13.71	0.33		X
	WI	5	126,852	13.19	0.46	X	
<b>Low-Duration States</b>	AR	6	67,191	12.94	0.25		
	OK	6	60,759	14.36	0.22		X
	LA	6	110,283	14.51	0.40	X	
	MO	7	193,860	14.52	0.71		X
	KS	7	60,004	14.53	0.22	X	
	MT	8	22,474	13.80	0.08		
	WY	8	10,047	14.13	0.04		
	HI	9	30,882	13.00	0.11		
	NV	9	53,816	14.60	0.20		X
<b>Low-Duration States</b>	AZ	9	91,442	14.71	0.33		
	NH	1	35,918	12.38	0.13		
	VA	3	237,954	12.33	0.87	X	
	AL	4	104,671	11.01	0.38		X
	SC	4	102,012	11.72	0.37		X
	GA	4	154,815	11.73	0.57	X	
	TN	4	184,164	12.93	0.67	X	
	IN	5	114,853	11.58	0.42		X
	NE	7	16,849	11.27	0.06		
<b>Low-Duration States</b>	IA	7	57,078	12.67	0.21		

TABLE B.1 (continued)

Selection Criteria	State	DOL Region	Number of EUC First Claims	Average Benefit Duration	Selection Probability	Supplementary Sample	
					23 States	States in Initial Sample	States in Supplementary Sample
	SD	8	3,560	10.66	0.01		
	UT	8	29,446	11.75	0.11		
	ND	8	14,681	12.17	0.05		X
	CO	8	66,902	12.47	0.24	X	
	ID	10	39,054	11.74	0.14		
<b>Total</b>			<b>9,215,995</b>	<b>15.80</b>	<b>23</b>	<b>23</b>	<b>12</b>

NOTE: EUC claims are for first payments based on regular UI, Unemployment Compensation for Ex-Servicemen (UCX), and Unemployment Compensation for Federal Employees (UCFE). The average benefit duration is for regular UI in 1991. It is computed from data in the U.S. Department of Labor, UI Database.

\*The weight is the state share of EUC claims times 51.

Finally, we decided to select, at a minimum, 10,000 EUC recipients and 10,000 UI-only recipients as the first stage of the sampling process. We intended to obtain administrative records for these samples and then select the smaller samples (2,500 EUC recipients and 1,500 UI-only recipients) for the interview. We chose 10,000 as the sample size for each of these administrative records samples to ensure that we had enough sample members to (1) complete 4,000 interviews on subsamples drawn from the latter three phases of EUC, and (2) examine the characteristics and experiences of EUC recipients by program phase.

## **B. IMPLEMENTATION OF THE SAMPLE DESIGN**

We implemented our sample design by contacting the 23 states we selected and asking them to select random samples of recipients who either collected EUC or who began collecting UI between January 1991 and September 1993. To reduce the burden on states, we did not ask them to give us separate EUC and UI-only samples. Instead, we asked for a single sample of recipients who met either criterion (collected EUC or collected UI during the relevant period). In addition, we used data on the number of EUC and UI first payments reported by states to the Unemployment Insurance Service to set sampling rates designed to meet our target of having a minimum of 10,000 EUC and 10,000 UI only sample members. Since the EUC population was smaller than the UI-only population, and since we were conservative in setting the sampling rates, this approach meant that we ended up with administrative records samples that were larger than our minimums.

Since our sample frame covered several years, we also asked states to provide administrative data on all benefit years established during this time frame by members of this sample. When we used administrative data for our analysis, we sometimes used the individual as the unit of analysis and we sometimes used the benefit year. However, we had to decide how to handle individuals with multiple benefit years in the interviewing subsample, since the interview used the benefit year begin

date to establish a time frame for the interview which began with the pre-benefits job. One option would have been to start with the earliest benefit year. We rejected that approach, however, because we felt that our main objective of representing the EUC population was better served by sampling benefit years for the interview. Hence, we assigned individuals to the EUC subsample if they ever collected EUC, and we began the interview with the benefit year that led directly to EUC. A few individuals had more than one EUC claim during Phases 3 through 5. In these cases, we randomly selected one of these claims as the start date for the interview. We also randomly selected a benefit year to start the interview for UI-only sample members with more than one benefit year.

A relatively large number of the states we selected were not able to participate in the study because of constraints on their programming resources or for other reasons. Specifically, 10 of the 23 states we initially contacted did not participate in the study. We addressed this situation by selecting a further random sample of 12 noncertainty states; of these, 7 did not participate. We ended up with samples from 18 states.

We encountered two further difficulties in implementing our design. First, two of the states that provided samples provided them too late for inclusion in the survey. Therefore, our survey sample is drawn from 16 states. Second, because we had difficulty locating sample members (as discussed more fully in Appendix C) our survey sample is smaller than planned and not distributed by state in the same proportions as planned.

Table B.2 reports final sample sizes, by state and by sample type. Our final sample included 28,420 individuals (34,484 benefit years) for whom we collected administrative data. It also included 1,341 EUC and 963 UI-only individuals for whom we collected survey data.

TABLE B.2  
EUC STUDY SAMPLE SIZES

States	Administrative Records Sample		Survey Samples	
	Individuals	Benefit Years	EUC	UI-only
High-Duration				
California	4,945	5,773	141	146
Connecticut	2,313	2,612	n.a.	n.a.
District of Columbia	521	581	38	19
Florida	1,566	1,840	87	42
Illinois	1,546	1,917	68	34
Maine	443	580	67	37
Minnesota	1,683	2,203	121	92
New Jersey	3,393	4,423	76	27
Pennsylvania	2,098	2,733	144	84
Texas	1,406	1,619	137	37
West Virginia	1,201	1,514	77	70
Medium-Duration States				
Kentucky	1,534	1,691	93	71
Louisiana	664	902	n.a.	n.a.
Oklahoma	893	1,013	63	40
Wisconsin	1,781	2,111	86	123
Low-Duration States				
Georgia	944	1,130	38	73
North Carolina	984	1,227	59	26
Tennessee	505	615	46	42
<b>Total</b>	<b>28,420</b>	<b>34,484</b>	<b>1,341</b>	<b>963</b>

n.a. = not applicable--sample received too late for inclusion in survey.

## C. WEIGHTS

We constructed weights for the administrative records and survey samples to produce nationally representative estimates. For the administrative records sample, the weights were designed to produce national estimates of the population of individuals receiving UI and/or EUC during the EUC period and national estimates of the benefit years established during that period. For the survey samples, the weights were designed to produce national estimates of the EUC Phase 3 through 5 population and national estimates of the UI-only population receiving UI during that period.

The major problem we faced in constructing these weights was that we had fewer (and, sometimes, different) states in the final sample than planned.<sup>9</sup> We addressed this problem by using external data on state-level UI and EUC activities reported by states to the Unemployment Insurance Service to compute national estimates of key EUC and UI program outcomes. In making these estimates, we treated each program separately; however, since most individuals who collected EUC also collected UI, we also computed estimates of key outcomes for the combined population (that is, individuals who collected under either program). We used data from our records samples to estimate the proportion of EUC recipients who did not begin collecting UI between January 1991 and September 1993.<sup>10</sup> We used the proportion for each state in our sample and the average for other states to compute the number of EUC first payments to individuals who did not collect UI. We then added this number to the number of UI first payments to compute the number of EUC and/or UI first payments made during our observation period. This unduplicated count of first payments was

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<sup>9</sup>The distribution of sample members by state was also different than planned but this did not present a major problem. We had random samples of recipients in each state in our records samples, and, although there was some nonresponse to the survey, we treated the survey samples in each state as simple random samples when constructing weights. We examine the appropriateness of this assumption in Appendix C.

<sup>10</sup>These are the reachback and EUC option claims.



divided into the number of weeks compensated and total payments under the two programs to produce our national EUC/UI estimates.

Using these estimates of national figures, we examined two alternative ways of weighting the state samples. Under the first alternative, we weighted the 18 states in the records sample to represent themselves. That is, we assigned weights such that the California sample represented California, the Connecticut sample represented Connecticut, and so on. This is a conservative approach that says that the sample only represents the 55 percent of the population found in the sampled states.

Under the second alternative, we weighted the 18 states in the records sample to represent the national population. We did this by grouping certainty and noncertainty states by stratum and adjusting the initial weight (the share of the total population represented by a state) assigned to each state to account for any nonresponse in the stratum. For example, we initially selected eight certainty and four noncertainty states in the high duration stratum but we ended up with six certainty and five noncertainty states whose weights, when summed, implied that this stratum equaled 54 percent of the EUC population. Since the high-duration states actually contained 65 percent of the EUC population, we increased each state weight to sum to 65 percent.

Our comparison of these weighting schemes (see Table B.3) indicated that either approach would produce estimates that appear close to our national estimates for the EUC population, the UI population, or the combined EUC/UI population. However, since the weights designed to represent the national population produced estimates closer to our national estimates, and since we would like to characterize our estimates as representing the nation, we chose to use the national weights in our analysis.

TABLE B.3  
COMPARISON OF ALTERNATIVE  
WEIGHTING SCHEMES

		Weighting Alternative	
	National Estimate	Weight States to Represent States in Sample	Weight States to Represent All States
EUC			
Average Weeks Compensated	17.6	18.5	17.5
Average Payments	\$3,080	\$3,152	\$2,916
Exhaustion Rate	54.5	57.9	55.3
UI			
Average Weeks Compensated	16.0	16.3	15.8
Average Payments	\$2,704	\$2,693	\$2,556
Exhaustion Rate	38.2	40.2	38.2
EUC/UI			
Average Weeks Compensated	21.0	21.6	20.6
Average Payments	\$3,620	\$3,599	\$3,373
Exhaustion Rate	N.A.	18.6	18.7

N.A. = not available.

We then computed weights for our records and survey samples designed to make these samples representative of the national populations of EUC and UI recipients. We created one weight for the records sample and two weights for the survey samples (see Table B.4). More specifically for the records sample, we created weights that when multiplied by the individuals or benefit years in the sample sum respectively to the total number of individuals who collected UI and /or EUC during the EUC period and that sum to the total number of benefit years established during this period. As noted previously, we defined the EUC period as including all individuals who received an EUC first payment and all individuals who received a UI first payment between January 1991 and September 1993 and did not collect EUC. As described above, we used data from our records samples to estimate the proportion of EUC first payments to recipients who did not also begin collecting UI between January 1991 and September 1993. We then used these figures to compute unduplicated counts of benefit years established during the EUC period. Finally, we used these numbers to adjust our initial sample weights by stratum, as described earlier. The resulting weights are applicable to individuals or benefit years included in our sample.

We used an analogous procedure for the survey samples to create weights for the EUC and UI-only samples that sum to national totals of EUC recipients who began collecting EUC during Phases 3 through 5 and UI-only recipients who began receiving UI between January 1992 and September 1993.

### **C. DESIGN EFFECTS**

The standard errors produced by most statistical programs are computed under the assumption that the samples used to compute estimates are simple random samples of the population. However, as we discussed previously, these standard errors underestimate the true standard errors for estimates

TABLE B.4  
EUC STUDY SAMPLE WEIGHTS

States	Records Sample	Survey Samples	
		EUC	UI-only
High-Duration			
California	801	5,997	14,559
Connecticut	363	n.a.	n.a.
District of Columbia	1,630	6,281	19,443
Florida	647	4,359	8,841
Illinois	705	5,277	14,862
Maine	1,633	3,562	9,984
Minnesota	430	1,973	4,015
New Jersey	256	4,916	13,175
Pennsylvania	628	3,690	7,058
Texas	900	3,225	14,972
West Virginia	626	3,100	5,277
Medium-Duration States			
Kentucky	898	4,298	11,821
Louisiana	1,683	n.a.	n.a.
Oklahoma	1,498	6,344	20,982
Wisconsin	719	4,647	6,823
Low-Duration States			
Georgia	1,549	8,422	10,076
North Carolina	1,147	7,404	15,141
Tennessee	2,846	6,957	17,512

n.a. = not applicable--sample received too late for inclusion in survey.

made with our samples, since they are not simple random samples of the national population. Instead, our samples are clustered by state, and this clustering increases standard errors.

We examined the degree to which simple random sample standard errors should be increased to account for the sample design. We computed these design effects using the SUDAAN computer program, which was developed at the Research Triangle Institute.<sup>11</sup> This program uses Taylor Series approximations to compute estimated variances using standard formulas that relate the size of the design effect to the relative size of two variables: (1) the component of the variance due to variation within individual clusters in the survey design, and (2) the component of variance due to differences between clusters in the relevant underlying population characteristics.

Since we examine a number of characteristics of sample members, and since the size of the design effect varies by characteristic, we computed design effects for a number of variables. Table B.5 shows the results of this exercise. These estimates range from a low, negligible effect of 1.02 for the percent female in the EUC sample to a high of 2.32 for the percent white. This range is not surprising, since the proportion of the UI population that is female is unlikely to vary among states (clusters in our sample) as much as the proportion that is white. Other important variables, such as the mean weekly benefit amounts and mean weeks on UI and EUC, which are likely to vary by state given differences in state laws and economic conditions, have design effects in the mid to high end of this range.

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<sup>11</sup>We report design effects computed as the proportional change in the standard error due to the survey design as compared to the standard error that could be achieved by a simple random sample of the same size, although design effects are often reported as the proportional change in the variance due to the survey design.

TABLE B.5  
DESIGN EFFECTS: SURVEY SAMPLES

Characteristic	EUC Sample			UI-only Sample		
	Mean	Standard Error	Design Effect	Mean	Standard Error	Design Effect
Demographic Characteristics						
Percent Female	43.8	1.4	1.02	40.8	2.5	1.59
Percent White	69.7	2.9	2.32	74.0	3.3	2.31
Mean Age	40.6	3.4	1.03	38.4	4.0	1.05
Pre-UC Labor Market Characteristics						
Percent Pre-UC Job in Manufacturing	32.6	2.0	1.57	33.2	2.3	1.51
Percent Expect Recall	28.3	1.8	1.45	38.1	1.8	1.96
UC Experiences						
Mean Weekly Benefit Amount	\$180	3.7	2.02	\$177	5.0	2.28
Mean Weeks UI	22.0	.4	2.19	11.9	5.1	1.80
Mean Weeks EUC	14.1	.3	1.31	n.a.	n.a.	n.a.
Percent Exhausted EUC	60.0	2.1	1.49	n.a.	n.a.	n.a.
Labor Market Outcomes (if Reemployed)						
Mean Months Until First Job	13.2	.4	1.12	6.9	.4	1.07
Ratio of First Job Weekly Wage to Pre-UC Weekly Wage	.9	.02	1.04	1.0	.02	1.02
	1,341			963		

SOURCE: Emergency Unemployment Compensation Survey.

NOTE: The design effects were computed using the SUDAAN program developed by Research Triangle Institute. While design effects are usually shown as the effect of the simple design on the variance, we report the effects on the standard error of the estimate.

Given this range in design effects, we chose to use the average (1.6) as a rough design effect to apply to our survey results. That is, when making comparisons between the EUC and UI samples, we inflated standard errors by 1.6 when determining which differences were statistically significant.

Table B.6 provides standard errors for the survey samples for binary variables used to estimate the prevalence of characteristics that can be expressed as a proportion or percent (for example, the percent expecting recall). These standard errors can be used to compute confidence intervals for such characteristics or to compute standard errors for difference of means tests. For example, we reported in Table III.6 that 23 percent of the UI-and-EUC sample expected recall by their pre-UI employer as compared to 49 percent for the EUC-only sample. Two-tailed 95 percent confidence intervals for these estimates would equal 1.96 times the appropriate standard error from Table III.6, which would be approximately +/-4 percent for the UI-and-EUC sample. The t-statistic for a difference of means test equals the difference between a characteristic for two groups divided by the standard error of the difference of means, which equals the square root of the sum of the variances of the two estimates. For example, the t-statistic for the difference in the expected recall rate for the UI-and-EUC and the EUC-only sample is  $5.1 \left[ (49-23)/\sqrt{(4.6^2+2.3^2)} \right]$ . This level indicates that the difference is statistically significant at the 99 percent confidence level for a two-tailed test.

TABLE B.6  
STANDARD ERRORS FOR PERCENTS  
EUC SURVEY

Characteristic Percent	EUC-UI	EUC-Only	UI-Only
50	2.5	4.6	2.6
40 (60)	2.4	4.5	2.5
30 (70)	2.3	4.2	2.4
20 (80)	2.0	3.7	2.1
10 (90)	1.5	2.8	1.5

NOTE: The standard errors were computed using the formula  $1.6 \sqrt{p(1-p)/n}$  where p is the percent of the population with a characteristic, and n is the sample size for the EUC-UI, EUC-only, or UI-only sample. The 1.6 factor is used to inflate the standard errors to account for design effects.



## **APPENDIX C**

### **SURVEY RESULTS AND NONRESPONSE BIAS ANALYSIS**



The EUC study design called for the selection of nationally representative samples of Unemployment Compensation (UC) recipients and the collection of Unemployment Insurance (UI) program data and, for a subsample, telephone survey data on the pre-layoff characteristics of recipients and their post-layoff labor market experiences. We implemented this design using a two-step process involving the random selection of states and the random selection of UC recipients in those states. Initially we selected 23 states for the sample, but, as discussed in Appendix A, not all states agreed to participate. In the end, 18 states provided data, with 16 doing so in time for inclusion in the survey. We then selected subsamples of EUC recipients and UI recipients who did not collect EUC (called the "UI-only sample") for the telephone survey. The EUC sample was chosen to represent individuals who began receiving EUC in July 1992 or later (that is, in EUC, Phases 3 through 5). The UI-only sample was chosen to represent individuals who began receiving UI during the period January 1992 through September 1993. These individuals would have collected EUC during the same time period as the EUC sample if they had continued onto EUC. The survey subsamples were restricted in this way to help minimize recall error. Even with this restriction, however, the recall period was long. The interviews were conducted between April 1996 and April 1997, which, on average, was three and a half years after the respondents' UC first payments.

This appendix provides information on the survey results, the number of completions, their distribution by state, and the reasons for nonresponse. It uses administrative records data to examine nonresponse and to assess the likelihood that survey results could be biased because of nonresponse to the survey.<sup>1</sup>

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<sup>1</sup>Another source of nonresponse that could affect our findings is nonresponse among the states selected for the survey. We address that issue in Appendix A.

## A. SURVEY RESULTS

We attempted interviews with 4,781 sample members and completed interviews with 2,304, yielding an overall response rate of 48 percent (Table C.1). This response rate varied slightly by sample; it was 46 percent for the EUC sample and 52 percent for the UI-only sample. It also varied by state as has been our experience in other, similar studies. It was highest in Minnesota and Wisconsin (just over 60 percent) and lowest in California and Texas (40 percent).

The overall response rate and the rates in each state were low, both in an absolute sense and in comparison to the rates achieved in prior surveys of UI recipients. For example, Corson and Dynarski (1990) report an overall response rate of 60 percent in their study of UI exhaustees. Response rates for states included in both studies were as much as 20 percentage points higher in the earlier survey.

Several reasons exist for the low response rate achieved in this study. The most important one is that it was difficult to locate sample members. As Table C.2 shows, 32 percent of the cases were not located (60 percent of the nonrespondents). The interview was conducted approximately three and a half years after the UC first payment was made, and the addresses and telephone numbers available from UC records were old. Having old, out-of-date addresses contributed to the difficulty we encountered in locating sample members; but, interestingly, 31 percent of the sample members in the exhaustee study also were not located, and the addresses in that study were, on average, only 20 months old.

Another factor contributing to the low response rate is that some individuals either did not complete the full interview (2 percent) or refused to be interviewed (10 percent). In this case, the experience in the exhaustee study was different; there were virtually no partial completes and the refusal rate was half that of this study. One possible reason for this difference is that this interview

TABLE C.1

## EUC SURVEY: NUMBER OF COMPLETES AND COMPLETION RATES, BY STATE

State	EUC		UI-Only		Total	
	Number	Percent	Number	Percent	Number	Percent
California	141	39.6	146	40.6	287	40.1
District of Columbia	38	38.4	19	61.3	57	43.8
Florida	87	38.5	42	48.8	129	41.3
Georgia	38	38.0	73	49.3	111	44.8
Illinois	68	42.5	34	42.5	102	42.5
Kentucky	93	51.7	71	46.7	164	49.4
Maine	67	52.8	37	62.7	104	55.9
Minnesota	121	63.0	92	63.0	213	63.0
North Carolina	59	49.6	26	53.1	85	50.6
New Jersey	76	40.6	27	46.6	103	42.0
Oklahoma	63	41.4	40	47.6	103	43.6
Pennsylvania	144	53.3	84	56.4	228	54.4
Tennessee	46	46.5	42	59.2	88	51.8
Texas	137	38.5	37	45.1	174	39.7
Wisconsin	86	62.8	123	61.5	209	62.0
West Virginia	77	50.7	70	61.4	147	55.3
<b>Total</b>	<b>1,341</b>	<b>46.1</b>	<b>963</b>	<b>51.5</b>	<b>2,304</b>	<b>48.5</b>

**TABLE C.2**  
**EUC SURVEY OUTCOMES**

<b>Interview Outcome</b>	<b>Percent</b>
Completion	48.2
Partial Completion	2.3
Refusal	10.0
Could Not Locate	31.8
Case Retired	5.7
Other	2.0
Total	100.0
<b>Total Cases Released</b>	<b>4,781</b>

was longer (it averaged 45 minutes) than the earlier interview (which averaged about 30 minutes). The interview length contributed to the partial completes, since some individuals refused to continue with the interview. Length may also have contributed to the refusals, since individuals were told approximately how long the interview was when they were asked to participate.

Finally, some cases (about six percent) were retired because we made multiple phone calls without reaching the potential respondent or for other reasons such as ill health, language barriers, or death (two percent).

## **B. POTENTIAL NONRESPONSE BIAS**

Results of the survey could be affected by nonresponse bias, particularly since the overall completion rate was quite low. If nonrespondents differ from respondents in a systematic way, inferences drawn from the interview data on the characteristics and labor market experiences of respondents could be misleading and not representative of the universe of UC recipients.

To analyze the implications of survey nonresponse for the analysis, we used UC administrative data that were available for both respondents and nonrespondents to explore differences in the baseline characteristics of respondents and nonrespondents and in UC outcomes. To perform this analysis, we used the weights described in Appendix A to create estimates for the respondent sample, which can be characterized as nationally representative of the UC population. We created comparable weights for nonrespondents, so that the weighted distribution of nonrespondents matched the weighted distribution of respondents by state and UC status (EUC and UI-only). This step was necessary because response rates differed by state and by UC status.

Our analysis shows (Table C.3) that there were a number of statistically significant differences between respondents and nonrespondents. Survey respondents were more likely than nonrespondents to be female, older, and nonminority. Respondents also had higher base period

TABLE C.3

## DIFFERENCES BETWEEN SURVEY RESPONDENTS AND NONRESPONDENTS

	Survey Respondents	Survey Nonrespondents	Total
<b>Pre-Layoff Characteristics</b>			
Female (Percent)	41.7	37.9**	39.8
Mean Age	39.3	37.2***	38.3
Race/Ethnicity <sup>a</sup>			
Caucasian (percent)	74.1	67.8***	71.0
African American (percent)	12.2	15.5***	13.9
Hispanic (percent)	9.8	12.2***	11.0
Other race/ethnicity (percent)	3.9	4.4***	4.1
In Manufacturing (Percent)	29.6	28.4	29.0
Base Period Earnings	18,568	16,568***	17,581
<b>UI and EUC Experience</b>			
UI Maximum Benefit Amount (Dollars)	4,347	4,026***	4,187
UI Weekly Benefit Amount (Dollars)	179	169***	174
Weeks of Potential UI Duration	24.0	23.5***	23.8
UI Collected (Dollars)	2,768	2,578**	2,673
Weeks UI Collected	15.4	14.9	15.1
Exhausted UI (Percent)	40.5	42.5	41.5
EUC Maximum Benefit Amount (Dollars)	3,357	3,296	3,326
EUC Weekly Benefit Amount (Dollars)	180	176	178
Weeks of Potential EUC Duration	18.6	18.6	18.6



TABLE C.3 (continued)

	Survey Respondents	Survey Nonrespondents	Total
EUC Collected (Dollars)	2,574	2,547	2,560
Weeks EUC Collected	14.1	14.5	14.3
Exhausted EUC (Percent)	21.8	23.0	22.4
Percent of EUC Claimants Who Collected EUC First <sup>b</sup>	5.0	4.8	5.9
<b>Unweighted Sample Size</b>	<b>2,304</b>	<b>2,477</b>	<b>4,781</b>

SOURCE: Weighted administrative records and survey data.

NOTE: Statistics for either the UI or the EUC program pertain only to those claimants who participated in that program.

<sup>a</sup>A chi-squared statistic was used to test the hypothesis that the racial distribution of survey respondents is the same as the distribution of survey nonrespondents.

<sup>b</sup>Claimants collecting EUC first include both claimants who subsequently collected UI and those that did not.

\*Significantly different from zero at the .10 level, two-tailed test.

\*\*Significantly different from zero at the .05 level, two-tailed test.

\*\*\*Significantly different from zero at the .01 level, two-tailed test.

earnings--hence, higher average weekly benefit amounts and entitlements for the regular UI program. They also collected more dollars of UI, but differences in other UI outcomes (weeks collected and the exhaustion rate) were not statistically significant. This pattern of differences between respondents and nonrespondents suggests that the respondent sample represented an older, more stable population than the nonrespondent sample--which is not surprising, given that the main reason for nonresponse was an inability to locate a sample member.

Although we find statistically significant differences between respondents and nonrespondents, we think that the broad conclusions drawn from the survey data in this report are not affected substantially by nonresponse. There are two reasons for this conclusion. First, the main focus of this report is on describing EUC experiences. While we found some differences in UI program entitlements and collections, we did not find statistically significant differences for EUC program variables. Respondents and nonrespondents had similar EUC weekly benefit amounts and entitlements, and there were no significant differences in EUC outcomes--dollars collected, weeks collected, exhaustion rate, or likelihood of choosing the option to collect EUC instead of UI.

Second, most of the differences we found are small (although statistically significant). For example, the respondent-nonrespondent difference in UI potential weeks is one-half week, and the difference in the UI weekly benefit amount is \$10. Because the completion rate was roughly 50 percent, the nonresponse adjusted estimate differs from the survey estimate by half these amounts. Similarly, the differences in baseline characteristics seem small. For example, mean age differs from the nonresponse adjusted estimate by one year, and the percent female differs by two percentage points. Even the estimates for race/ethnicity differ at most by three percentage points. These differences seem small; as stated earlier, we think the broad conclusions we reach using the survey data are unaffected by nonresponse.

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